

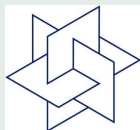
Scenario Generation in Stochastic Programming with Application to Optimizing Electricity Portfolios under Uncertainty

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Stochastic programming and approximation issues

We consider a stochastic program of the form

$$\min \left\{ \int_{\Xi} \Phi(x, \xi) P(d\xi) : x \in X \right\},$$

where $X \subseteq \mathbb{R}^m$ is a constraint set, P a probability distribution on $\Xi \subseteq \mathbb{R}^d$, and $f := \Phi(x, \cdot)$ is a **decision-dependent integrand**.

Any approach to solving such models computationally requires to replace the integral by a **quadrature rule**

$$Q_{n,d}(f) = \sum_{i=1}^n w_i f(\xi^i),$$

with weights $w_i \in \mathbb{R}$ and scenarios $\xi^i \in \Xi$, $i = 1, \dots, n$.

If the natural condition $w_i \geq 0$ and $\sum_{i=1}^n w_i = 1$ is satisfied, $Q_{n,d}(f)$ allows the interpretation as integral with respect to the **discrete probability measure** Q_n having scenarios ξ^i with probabilities w_i , $i = 1, \dots, n$.

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Example 1: Linear two-stage stochastic programs

We consider two-stage linear stochastic programs:

$$\min \left\{ \langle c, x \rangle + \int_{\Xi} \varphi(q(\xi), h(\xi) - Tx) P(d\xi) : x \in X \right\}$$

where $c \in \mathbb{R}^m$, X is a convex polyhedral subset of \mathbb{R}^m , Ξ a closed subset of \mathbb{R}^d , T a (r, m) -matrix, $h(\cdot)$ and $q(\cdot)$ are affine mappings on \mathbb{R}^d , P a Borel probability measure on Ξ and

$$\begin{aligned} \varphi(q, t) &= \inf \{ \langle q, y \rangle : Wy = t, y \geq 0 \} \\ &= \sup \{ \langle t, z \rangle : W^\top z \leq q \} \end{aligned}$$

where $q \in \mathbb{R}^{\bar{m}}$, W a (r, \bar{m}) -matrix (having rank r) and t varies in the polyhedral cone $W(\mathbb{R}_+^{\bar{m}})$. There exist matrices C_j and polyhedral cones \mathcal{K}_j , $j = 1, \dots, \ell$, decomposing $\text{dom } \varphi$ such that $\varphi(q, t) = \langle C_j q, t \rangle$, $\forall (q, t) \in \mathcal{K}_j$. Hence, the integrand is

$$\Phi(x, \xi) = \langle c, x \rangle + \max_{j=1, \dots, \ell} \langle C_j q(\xi), h(\xi) - Tx \rangle$$

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Example 2: Linear multi-stage stochastic programs

Let $\{\xi_t\}_{t=1}^T$ be a discrete-time stochastic data process defined on some probability space $(\Omega, \mathcal{F}, \mathbb{P})$ and with ξ_1 deterministic. The stochastic decision x_t at period t is assumed to be measurable with respect to $\mathcal{F}_t := \sigma(\xi_1, \dots, \xi_t)$ (**nonanticipativity**).

$$\min \left\{ \mathbb{E} \left(\sum_{t=1}^T \langle b_t(\xi_t), x_t \rangle \right) \left| \begin{array}{l} x_t \in X_t, x_t \in L_p(\Omega, \mathcal{F}_t, \mathbb{P}; \mathbb{R}^{m_t}) \\ \sum_{\tau=0}^{t-1} A_{t,\tau} x_{t-\tau} = h_t(\xi_t) \\ (t = 1, \dots, T) \end{array} \right. \right\}$$

where the sets X_t are convex polyhedral in \mathbb{R}^{m_t} , $A_{t,\tau}$, $\tau = 0, \dots, t-1$, are matrices and the vectors $b_t(\cdot)$ and $h_t(\cdot)$ depend affine linearly on ξ_t , $t = 1, \dots, T$.

The integrand $\Phi = \Phi_1$ is given by dynamic programming

$$\Phi_{t-1}(x^{t-1}, \xi^t) = \inf_{x_t \in X_t} \left\{ \langle b_t(\xi_t), x_t \rangle + \mathbb{E}(\Phi_t(x^t, \xi^{t+1}) | \mathcal{F}_t) \left| \begin{array}{l} \sum_{\tau=0}^{t-1} A_{t,\tau} x_{t-\tau} = h_t(\xi_t) \end{array} \right. \right\},$$

where $t = 2, \dots, T$, $\Phi_T \equiv 0$, $x^t = (x_1, \dots, x_t)$, $\xi^t = (\xi_1, \dots, \xi_t)$.

Assumption: P has a density ρ w.r.t. λ^d .

Now, we set $\mathcal{F} = \{\Phi(\cdot, x)\rho(\cdot) : x \in X\}$ and assume that the set \mathcal{F} is a bounded subset of some linear normed space F_d with norm $\|\cdot\|_d$ and unit ball $\mathbb{B}_d = \{f \in F_d : \|f\|_d \leq 1\}$.

The absolute error of the quadrature rule $Q_{n,d}$ is

$$e(Q_{n,d}) = \sup_{f \in \mathbb{B}_d} \left| \int_{\Xi} f(\xi) d\xi - \sum_{i=1}^n w_i f(\xi^i) \right|$$

and the approximation criterion is based on the relative error and a given tolerance $\varepsilon > 0$, namely, it consists in finding the smallest number $n_{\min}(\varepsilon, Q_{n,d}) \in \mathbb{N}$ such that

$$e(Q_{n,d}) \leq \varepsilon e(Q_{0,d}),$$

holds, where $Q_{0,d}(f) = 0$ and, hence, $e(Q_{0,d}) = \|I_d\|$ with

$$I_d(f) = \int_{\Xi} f(\xi) d\xi.$$

The behavior of both quantities depends heavily on the normed space F_d and the set \mathcal{F} , respectively.

It is **desirable** that an estimate of the form

$$n_{\min}(\varepsilon, Q_{n,d}) \leq C d^q \varepsilon^{-p} \quad (\text{'tractability'})$$

is valid for some constants $q \geq 0$, $C, p > 0$ and for every $\varepsilon \in (0, 1)$. Of course, $q = 0$ is highly desirable for **high-dimensional problems**.

Proposition: (Stability)

Let the set X be compact. Then there exists $L > 0$ such that

$$\left| \inf_{x \in X} \int_{\Xi} \Phi(\xi, x) \rho(\xi) d\xi - \inf_{x \in X} \sum_{i=1}^n w_i \Phi(\xi^i, x) \rho(\xi^i) \right| \leq L e(Q_{n,d}).$$

The solution set mapping is outer semicontinuous at P .

Alternatively, we look for a suitable set \mathcal{F} of functions such that $\{C\Phi(\cdot, x) : x \in X\} \subseteq \mathcal{F}$ for some constant $C > 0$ and, hence,

$$e(Q_{n,d}) \leq \frac{1}{C} \sup_{f \in \mathcal{F}} \left| \int_{\Xi} f(\xi) P(d\xi) - \int_{\Xi} f(\xi) Q_n(d\xi) \right| = D(P, Q_n),$$

and that D is a metric distance between probability distributions.

Example: L_p -minimal metric ℓ_p (or Wasserstein metric) of order $p \geq 1$

$$\ell_p(P, Q) := \left(\inf \left\{ \int_{\Xi \times \Xi} \|\xi - \tilde{\xi}\|^p \eta(d\xi, d\tilde{\xi}) \mid \pi_1 \eta = P, \pi_2 \eta = Q \right\} \right)^{\frac{1}{p}}$$

It holds

$$\ell_p(P, Q) = \inf \{ \|\xi - \tilde{\xi}\|_p \mid \mathcal{L}(\xi) = P, \mathcal{L}(\tilde{\xi}) = Q \}$$

$$\ell_1(P, Q) = \sup \left\{ \left| \int_{\Xi} f(\xi) (P - Q)(d\xi) \right| : |f(\xi) - f(\tilde{\xi})| \leq \|\xi - \tilde{\xi}\| \right\}$$

by definition and duality, respectively.

Examples of normed spaces F_d :

- (a) The Banach space $F_d = \text{Lip}(\mathbb{R}^d)$ of Lipschitz continuous functions equipped with the norm

$$\|f\|_d = |f(0)| + \sup_{\xi \neq \tilde{\xi}} \frac{|f(\xi) - f(\tilde{\xi})|}{\|\xi - \tilde{\xi}\|}.$$

The best possible convergence rate is $e(Q_{n,d}) = O(n^{-\frac{1}{d}})$.

It is attained for $w_i = \frac{1}{n}$ and certain ξ^i , $i = 1, \dots, n$, if P has finite moments of order $1 + \delta$ for some $\delta > 0$. (Graf-Luschgy 00)

- (b) The tensor product Sobolev space

$$F_{d,\gamma} = \mathcal{W}_{2,\text{mix}}^{(1,\dots,1)}([0, 1]^d) = \bigotimes_{j=1}^d W_2^1([0, 1])$$

of real functions on $[0, 1]^d$ having first order mixed weak derivatives with the (weighted) norm

$$\|f\|_{d,\gamma} = \left(\sum_{u \subset D} \gamma_u^{-1} \int_{[0,1]^{|u|}} \left| \frac{\partial^{|u|}}{\partial \xi^u} f(\xi^u, 1^{-u}) \right|^2 d\xi^u \right)^{\frac{1}{2}},$$



where $D = \{1, \dots, d\}$, $\gamma_1 \geq \gamma_2 \geq \dots \geq \gamma_d > 0$, $\gamma_\emptyset = 1$ and

$$\gamma_u = \prod_{j \in u} \gamma_j \quad (u \subseteq D).$$

For n prime, $w_i = \frac{1}{n}$, and a suitable choice of (γ_j) , points $\xi^i \in [0, 1]^d$, $i = 1, \dots, n$, can be constructed such that

$$e(Q_{n,d}) \leq C(\delta) n^{-1+\delta} \|I_d\|$$

for some $C(\delta) > 0$ (not depending on d) and all $0 < \delta \leq \frac{1}{2}$.

(Sloan, Woźniakowski 98, Kuo 03)

Scenario generation methods

We will discuss the following three scenario generation methods for stochastic programs *without nonanticipativity constraints*:

- (a) [Monte Carlo sampling](#) from the underlying probability distribution P on \mathbb{R}^d (Shapiro 03).
- (b) [Optimal quantization of probability distributions](#) (Pflug-Pichler 10).
- (c) [Quasi-Monte Carlo methods](#) (Koivu-Pennanen 05, Homem-de-Mello 06).

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Monte Carlo sampling

Monte Carlo methods are based on drawing independent identically distributed (iid) Ξ -valued random samples $\xi^1(\cdot), \dots, \xi^n(\cdot), \dots$ (defined on some probability space $(\Omega, \mathcal{A}, \mathbb{P})$) from an underlying probability distribution P (on Ξ) such that

$$Q_{n,d}(\omega)(f) = \frac{1}{n} \sum_{i=1}^n f(\xi^i(\omega)),$$

i.e., $Q_{n,d}(\cdot)$ is a random functional, and it holds

$$\lim_{n \rightarrow \infty} Q_{n,d}(\omega)(f) = \int_{\Xi} f(\xi) P(d\xi) = \mathbb{E}(f) \quad \mathbb{P}\text{-almost surely}$$

for every real continuous and bounded function f on Ξ .

If P has finite moment of order $r \geq 1$, the error estimate

$$\mathbb{E} \left(\left| \frac{1}{n} \sum_{i=1}^n f(\xi^i(\omega)) - \mathbb{E}(f) \right|^r \right) \leq \frac{\mathbb{E}((f - \mathbb{E}(f))^r)}{n^{r-1}}$$

is valid.

Hence, the **mean square convergence rate** is

$$\|Q_{n,d}(\omega)(f) - \mathbb{E}(f)\|_{L_2} = \sigma(f)n^{-\frac{1}{2}},$$

where $\sigma^2(f) = \mathbb{E}((f - \mathbb{E}(f))^2)$.

The latter holds without any assumption on f except $\sigma(f) < \infty$.

Advantages:

- (i) MC sampling works *for (almost) all integrands*.
- (ii) The machinery of probability theory is available.
- (iii) The convergence *rate does not depend on d* .

Deficiencies: (Niederreiter 92)

- (i) There exist 'only' *probabilistic error bounds*.
- (ii) Possible regularity of the integrand *does not improve* the rate.
- (iii) Generating (independent) random samples is *difficult*.

Practically, iid samples are approximately obtained by **pseudo random number generators** as uniform samples in $[0, 1]^d$ and later transformed to more general sets Ξ and distributions P .

Excellent pseudo random number generator: [Mersenne Twister](#)

(Matsumoto-Nishimura 98).

Survey: L'Ecuyer 94.

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Optimal quantization of probability measures

Let D be a distance of probability measures on \mathbb{R}^d such that the underlying stochastic program behaves stable w.r.t. D (Römisch 03).

Example:

L_p -minimal metric ℓ_p for $p \geq 1$, i.e.

$$\ell_p(P, Q) = \inf\{(\mathbb{E}(\|\xi - \eta\|^p))^{\frac{1}{p}} : \mathcal{L}(\xi) = P, \mathcal{L}(\eta) = Q\}$$

Let P be a given probability distribution on \mathbb{R}^d . We are looking for a discrete probability measure Q_n with support

$$\text{supp}(Q_n) = \{\xi^1, \dots, \xi^n\} \quad \text{and} \quad Q_n(\{\xi^i\}) = \frac{1}{n}, \quad i = 1, \dots, n,$$

that is the **best approximation to P with respect to D** , i.e.,

$$D(P, Q_n) = \min\{D(P, Q) : |\text{supp}(Q)| = n, Q \text{ is uniform}\}.$$

Existence of best approximations, called **optimal quantizers**, and their best possible convergence rate $O(n^{-\frac{1}{d}})$ is well known for ℓ_p (Graf-Luschgy 00).

However, in general, the function

$$\Psi_D(\xi^1, \dots, \xi^n) := D\left(P, \frac{1}{n} \sum_{i=1}^n \delta_{\xi^i}\right)$$

and, in particular,

$$\Psi_{\ell_p}(\xi^1, \dots, \xi^n) = \left(\int_{\mathbb{R}^d} \min_{i=1, \dots, n} \|\xi - \xi^i\|^p P(d\xi) \right)^{\frac{1}{p}}$$

is **nonconvex and nondifferentiable** on \mathbb{R}^{dn} .

Hence, the global minimization of Ψ_D is not an easy task.

Algorithmic procedures for minimizing Ψ_{ℓ_r} globally may be based on **stochastic gradient (type) algorithms, stochastic approximation methods and stochastic branch-and-bound techniques** (e.g. Pflug 01, Hochreiter-Pflug 07, Pagés 97, Pagés et al 04).

However, **asymptotically optimal quantizers can be determined explicitly** in a number of cases (Pflug-Pichler 10).

Quasi-Monte Carlo methods

The idea of Quasi-Monte Carlo (QMC) methods is to replace random samples in Monte Carlo methods by deterministic points ξ^i , $i \in \mathbb{N}$, that are **uniformly distributed** in $[0, 1]^d$. QMC is of the form

$$Q_{n,d}(f) = \frac{1}{n} \sum_{i=1}^n f(\xi^i)$$

The uniform distribution property may be defined in terms of the so-called **star-discrepancy** of ξ^1, \dots, ξ^n

$$D_n^*(\xi^1, \dots, \xi^n) := \sup_{\xi \in [0,1]^d} \left| \lambda^d([0, \xi)) - \frac{1}{n} \sum_{i=1}^n \mathbf{1}_{[0, \xi)}(\xi^i) \right|,$$

by calling a sequence $(\xi^i)_{i \in \mathbb{N}}$ **uniformly distributed** in $[0, 1]^d$ if

$$D_n^*(\xi^1, \dots, \xi^n) \rightarrow 0 \quad \text{for } n \rightarrow \infty.$$

A **classical result** due to Roth 54 states

$$D_n^*(\xi^1, \dots, \xi^n) \geq B_d \frac{(\log n)^{\frac{d-1}{2}}}{n}$$

for some constant B_d and all sequences (ξ^i) in $[0, 1]^d$.

Classical convergence results:

Theorem: (Proinov 88)

If the real function f is continuous on $[0, 1]^d$, then there exists $C > 0$ such that

$$|Q_{n,d}(f) - I_d(f)| \leq C\omega_f\left(D_n^*(\xi^1, \dots, \xi^n)^{\frac{1}{d}}\right),$$

where $\omega_f(\delta) = \sup\{|f(\xi) - f(\tilde{\xi})| : \|\xi - \tilde{\xi}\| \leq \delta, \xi, \tilde{\xi} \in [0, 1]^d\}$ is the modulus of continuity of f .

Theorem: (Koksma-Hlawka 61)

If f is of bounded variation in the sense of Hardy and Krause, it holds

$$|I_d(f) - Q_{n,d}(f)| \leq V_{\text{HK}}(f)D_n^*(\xi^1, \dots, \xi^n).$$

for any $n \in \mathbb{N}$ and any $\xi^1, \dots, \xi^n \in [0, 1]^d$.

There exist sequences (ξ^i) in $[0, 1]^d$ such that

$$D_n^*(\xi^1, \dots, \xi^n) = O(n^{-1}(\log n)^{d-1}).$$

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First general construction: **Nets** (Sobol 69, Niederreiter 87)

Elementary subintervals E of $[0, 1)^d$ in base b :

$$E = \prod_{j=1}^d \left[\frac{a_j}{b^{d_j}}, \frac{a_j + 1}{b^{d_j}} \right),$$

with $a_i, d_i \in \mathbb{Z}_+, 0 \leq a_i < d_i, i = 1, \dots, d$.

A set of b^m points in $[0, 1)^d$ is a **(t, m, d) -net in base b** if every elementary subinterval E in base b with $\lambda^d(E) = b^{t-m}$ contains b^t points ($m, t \in \mathbb{Z}_+, m > t$).

A sequence (ξ^i) in $[0, 1)^d$ is a **(t, d) -sequence in base b** if, for all integers $k \in \mathbb{Z}_+$ and $m > t$, the set

$$\{\xi^i : kb^m \leq i < (k+1)b^m\}$$

is a (t, m, d) -net in base b .

Proposition: $(0, d)$ -sequences **exist** if $d \leq b$.

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Theorem: A $(0, m, d)$ -net $\{\xi^i\}$ in base b satisfies

$$D_n^*(\xi^i) \leq A_d(b) \frac{(\log n)^{d-1}}{n} + O\left(\frac{(\log n)^{d-2}}{n}\right).$$

with reasonably small constants $A_d(b)$.

Special cases: Sobol, Faure and Niederreiter sequences.

Second general construction: Lattices (Korobov 59, Sloan-Joe 94)

Let $g \in \mathbb{Z}^d$ and consider the **lattice points**

$$\left\{ \xi^i = \left\{ \frac{i}{n} g \right\} : i = 1, \dots, n \right\},$$

where $\{z\}$ is defined componentwise and for $z \in \mathbb{R}_+$ it is the *fractional part* of z , i.e., $\{z\} = z - \lfloor z \rfloor \in [0, 1)$.

Randomly shifted lattice points with a uniform random vector Δ :

$$\left\{ \xi^i = \left\{ \frac{i}{n} g + \Delta \right\} : i = 1, \dots, n \right\},$$

There is a **component-by-component construction algorithm** for g such that for some constant $C(\delta)$ and all $0 < \delta \leq \frac{1}{2}$

$$e(Q_{n,d}) \leq C(\delta) n^{-1+\delta} \|I_d\| \quad (\text{Sloan-Kuo 05, Kuo 03}).$$

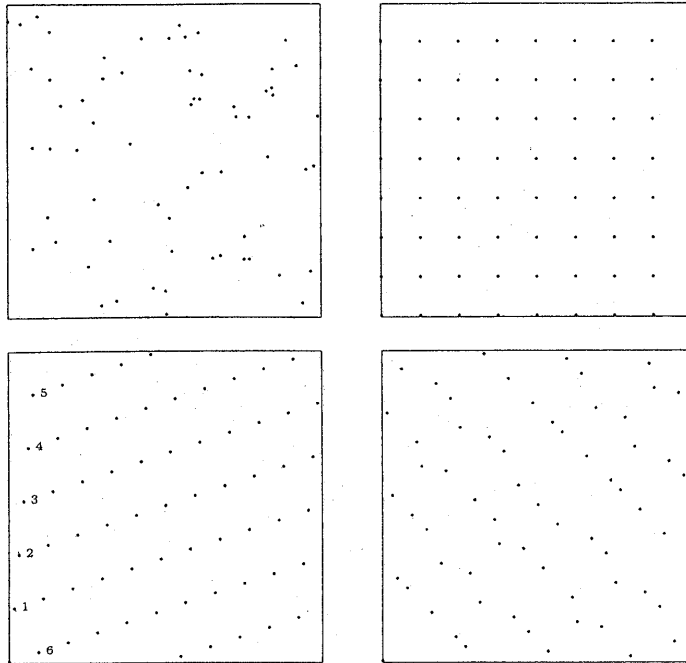


Fig. 5.3 Four different point sets with $n = 64$: random (top left), rectangular grid (top right), Korobov lattice (bottom left), and Sobol' (bottom right).

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Convergence rates for unbounded integrands ?

(Kuo-Sloan-Wasilkowski-Waterhouse 10)

Let us consider

$$I_{d,\rho}(f) = \int_{\mathbb{R}^d} f(\xi)\rho(\xi)d\xi \quad \text{with} \quad \rho(\xi) = \prod_{j=1}^d \phi(\xi_j)$$

and strictly positive ϕ (w.l.o.g.).

Transformation:

$$I_{d,\rho}(f) = I_d(g) = \int_{(0,1)^d} g(u)du, \quad \text{where}$$

$$g(u) = f(\Phi^{-1}(u)) := f(\Phi^{-1}(u_1), \dots, \Phi^{-1}(u_d)) \quad \text{and} \quad \Phi(\xi) = \int_{-\infty}^{\xi} \phi(t)dt$$

Absolute error:

$$e(Q_{n,d}) = \sup_{f \in \mathbb{B}_d} \left| \int_{(0,1)^d} f(\Phi^{-1}(u))du - \frac{1}{n} \sum_{i=1}^n f(\Phi^{-1}(u^i)) \right|$$

where $u^i \in (0,1)^d$, $i = 1, \dots, n$.

Rates of convergence for unbounded integrands are known for several densities ϕ and close to those for $[0,1]^d$.

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Is QMC efficient in stochastic programming ?

Problem: Typical integrands in linear stochastic programming are not smooth and, hence, do not belong to the relevant function spaces in general.

Idea: Study of the **efficient dimension** of typical integrands.

ANOVA-decomposition of f :

$$f = \sum_{u \subseteq D} f_u,$$

where $f_\emptyset = I_d(f) = I_D(f)$ and recursively

$$f_u = I_{-u}(f) + \sum_{v \subseteq u} (-1)^{|u|-|v|} I_{u-v}(I_{-u}(f)),$$

where I_{-u} means integration with respect to ξ_j in $[0, 1]$, $j \in D \setminus u$ and $D = \{1, \dots, d\}$. Hence, f_u is essentially as smooth as $I_{-u}(f)$ and does not depend on ξ^{-u} .

Proposition: The functions $\{f_u\}_{u \subseteq D}$ are **orthogonal** in $L_2([0, 1]^d)$.

We set $\sigma^2(f) = \|f - I_d(f)\|_{L_2}^2$ and have

$$\sigma^2(f) = \|f\|_{L_2}^2 - (I_d(f))^2 = \sum_{\emptyset \neq u \subseteq D} \|f_u\|_{L_2}^2.$$

The truncation dimension d_t of f is the smallest $d_t \in \mathbb{N}$ such that

$$\sum_{u \subseteq \{1, \dots, d_t\}} \|f_u\|_{L_2}^2 \geq \alpha \sigma^2(f) \quad (\text{where } \alpha \in (0, 1) \text{ is close to } 1).$$

Then

$$\|f - \sum_{u \subseteq \{1, \dots, d_t\}} f_u\|_{L_2}^2 \leq (1 - \alpha) \sigma^2(f).$$

Most of the ANOVA terms f_u may be smoother than f under certain conditions.

(Griebel-Kuo-Sloan 10).



A note on scenario reduction

Assume that the stochastic program behaves stable with respect to ℓ_p for some $p \geq 1$.

Let us consider discrete probability distributions P with scenarios ξ^i and probabilities p_i , $i = 1, \dots, N$, and Q being supported by a given subset of scenarios ξ^j , $j \in J \subset \{1, \dots, N\}$, of P .

The best approximation of P with respect to ℓ_p given an index set J exists and is denoted by Q^* . It has the distance

$$D_J := \ell_p(P, Q^*) = \min_Q \ell_p(P, Q) = \left(\sum_{i \in J} p_i \min_{j \notin J} \|\xi^i - \xi^j\|^p \right)^{\frac{1}{p}}$$

and the probabilities $q_j^* = p_j + \sum_{i \in J_j} p_i$, $\forall j \in J$, where

$J_j := \{i \in J : j = j(i)\}$ and $j(i) \in \arg \min_{j \notin J} \|\xi^i - \xi^j\|$, $\forall i \in J$

(optimal redistribution) (Dupačová-Gröwe-Römisch 03).

For **mixed-integer two-stage stochastic programs** the relevant distance is a **polyhedral discrepancy**. In that case, the new weights have to be determined by linear programming (Henrion-Küchler-Römisch 08, 09).

Determining the **optimal index set** J with prescribed cardinality $N - n$ is a **combinatorial optimization problem**:

$$\min \{D_J : J \subset \{1, \dots, N\}, |J| = N - n\}$$

Hence, the problem of finding the optimal index set J of scenarios to delete is **\mathcal{NP} -hard** and **polynomial time algorithms are not available in general**.

\implies **Heuristics are used to determine J .**

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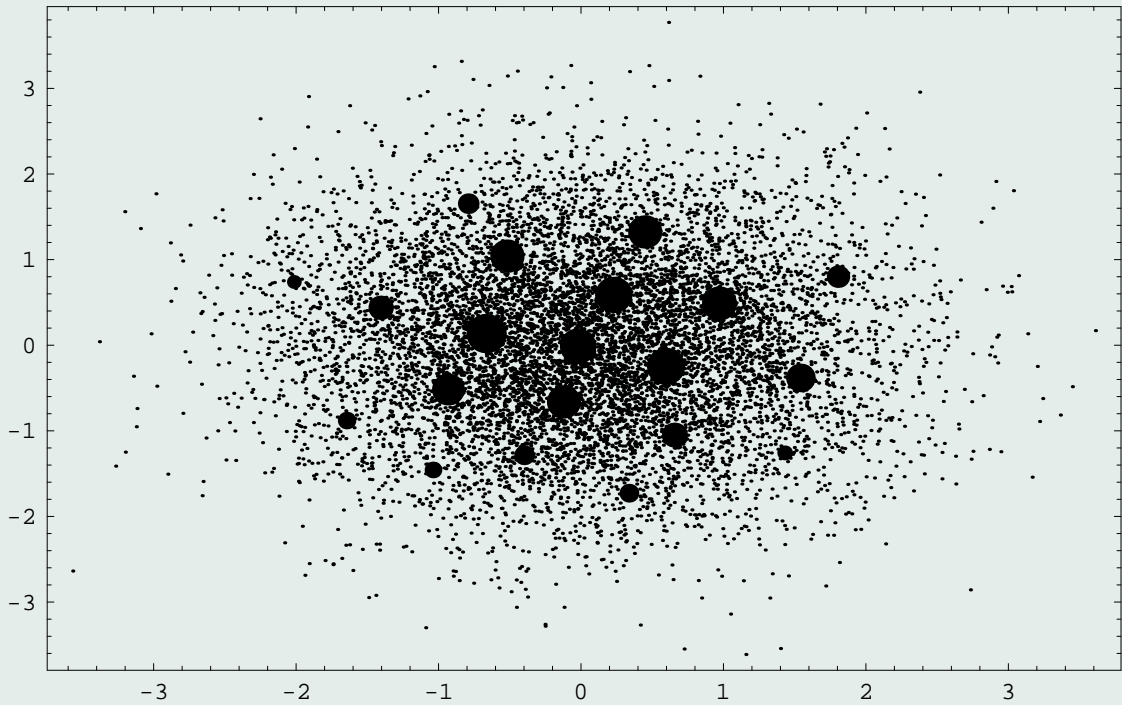
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Scenario reduction w.r.t. ℓ_1 from $N=10\,000$ MC samples of $N(0, I)$ in \mathbb{R}^2 to $n = 20$. The diameters of the circles are proportional to their probabilities

Generation of scenario trees

In **multistage stochastic programs** the decisions x have to satisfy the additional **information constraint** that x_t is measurable with respect to $\mathcal{F}_t = \sigma(\xi_\tau, \tau = 1, \dots, t)$, $t = 1, \dots, T$. The increase of the σ -fields \mathcal{F}_t w.r.t. t is reflected by approximating the underlying stochastic process $\xi = (\xi_t)_{t=1}^T$ by scenarios forming a **scenario tree**.

Some recent approaches:

- (1) **Bound-based approximation methods**: Kuhn 05, Casey-Sen 05.
- (2) **Monte Carlo-based schemes**: Shapiro 03, 06.
- (3) **Quasi-Monte Carlo methods**: Pennanen 06, 09 .
- (4) **Moment-matching principle**: Høyland-Kaut-Wallace 03.
- (5) **Optimal quantization**: Pagés et al. 03.
- (6) **Stability-based approximations**: Hochreiter-Pflug 07, Mirkov-Pflug 07, Pflug-Pichler 10, Heitsch-Römisch 05, 09.

Survey: Dupačová-Consigli-Wallace 00

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Theoretical basis of (6):

Quantitative stability results for multi-stage stochastic programs.

(Heitsch-Römisch-Strugarek 06; Mirkov-Pflug 07, Pflug 09)

Scenario tree generation: (Heitsch-Römisch 09)

- (i) Generate a number of **scenarios** by one of the methods discussed earlier.
- (ii) **Construction of a scenario tree** out of these scenarios by **recursive scenario reduction and bundling over time** such that the optimal value stays within a prescribed tolerance.

Implementation: GAMS-SCENRED 2.0 (developed by H. Heitsch)

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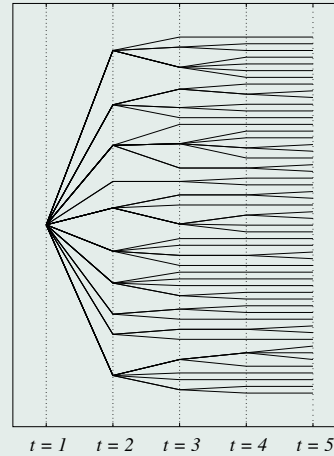
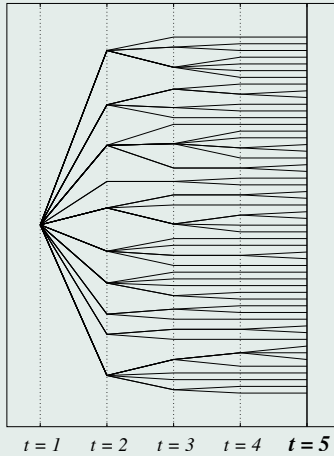
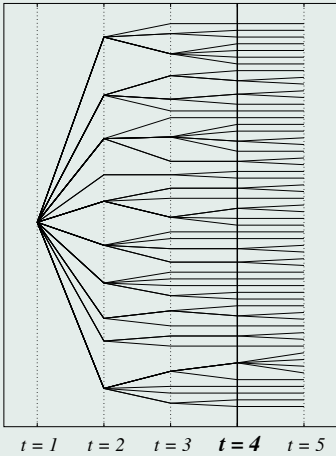
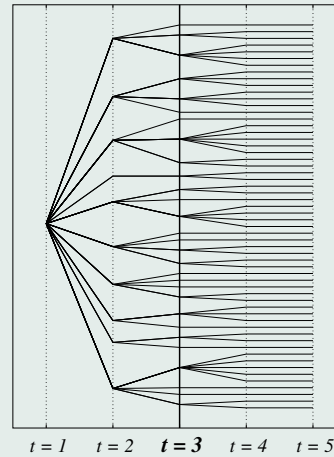
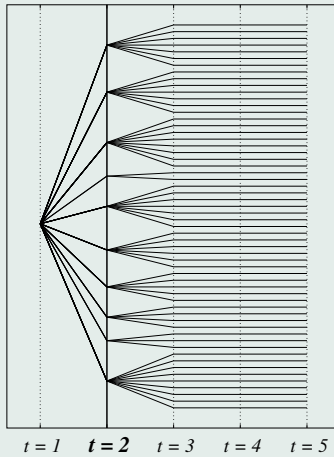
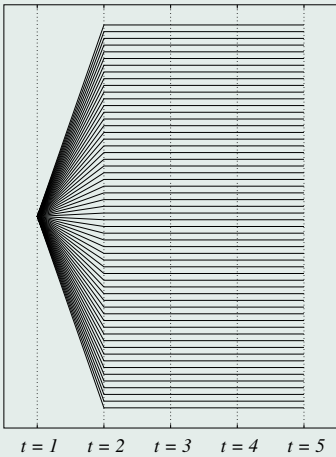


Illustration of the **forward tree generation** for an example including $T=5$ time periods starting with a scenario fan containing $N=58$ scenarios

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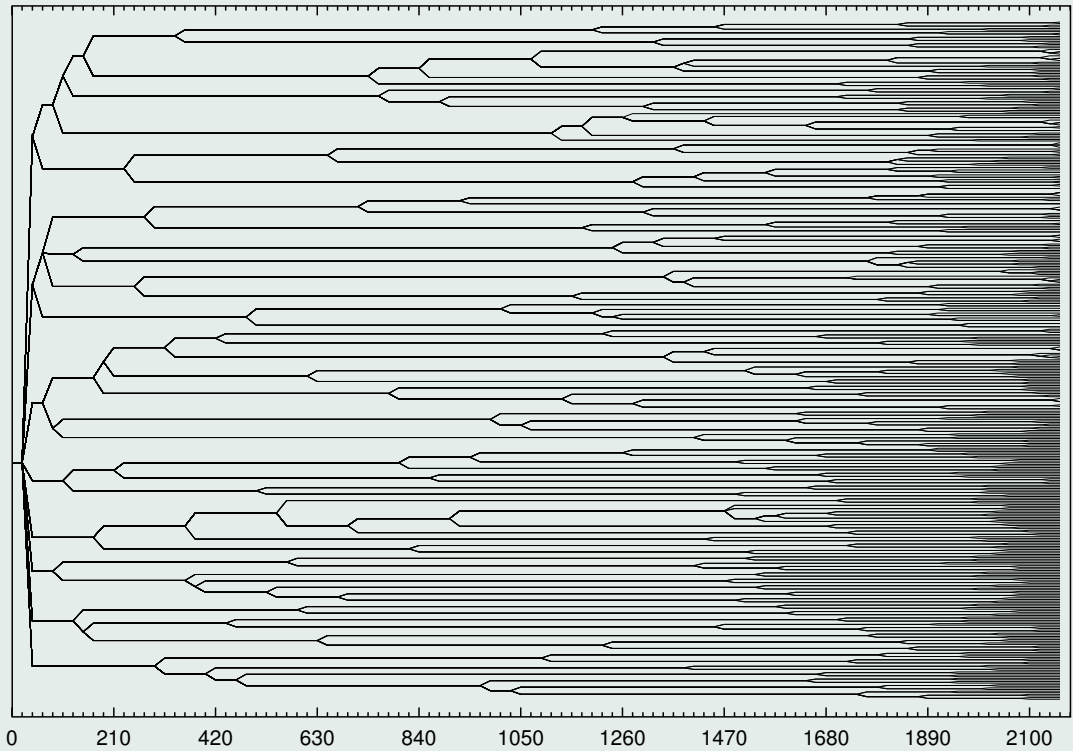
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Two-yearly demand-inflow scenario tree with weekly branchings for French EDF

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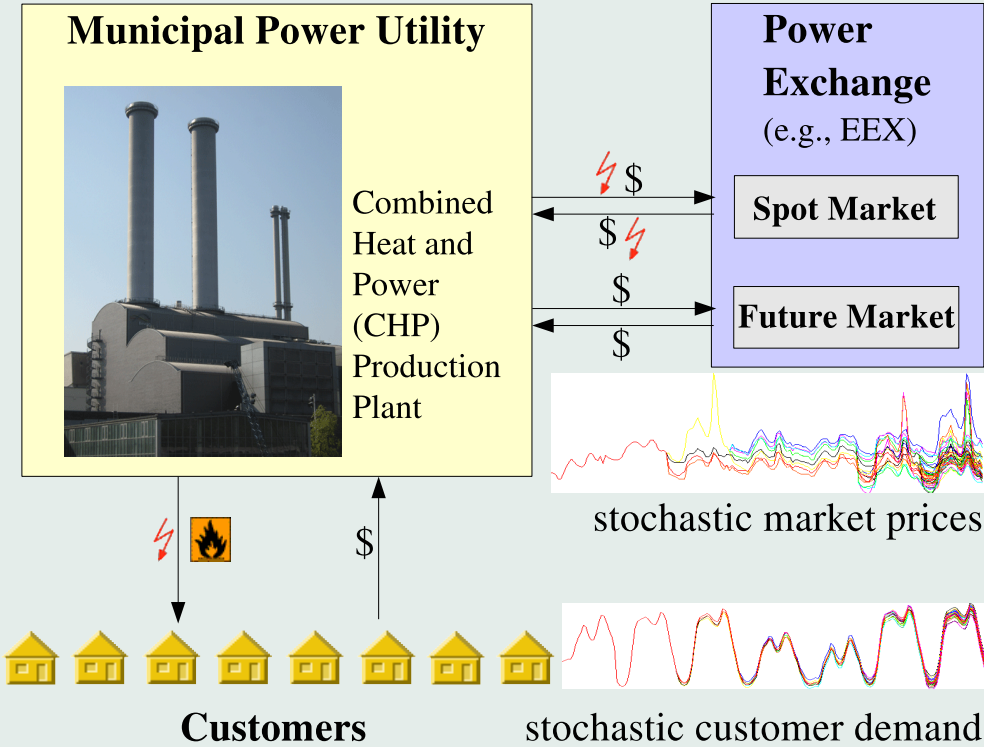
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Mean-Risk Electricity Portfolio Management



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We consider the [electricity portfolio management](#) of a German municipal [electric power company](#). Its portfolio consists of the following positions:

- [power production](#) (based on company-owned thermal units),
- [bilateral contracts](#),
- (physical) [\(day-ahead\) spot market trading](#) (e.g., [European Energy Exchange \(EEX\)](#)) and
- (financial) [trading of futures](#).

The time horizon is discretized into [hourly intervals](#). The underlying stochasticity consists in a [multivariate stochastic load and price process](#) that is approximately represented by a finite number of scenarios. The objective is to [maximize the total expected revenue and to minimize the risk](#). The portfolio management model is a large scale [\(mixed-integer\) multi-stage stochastic program](#).

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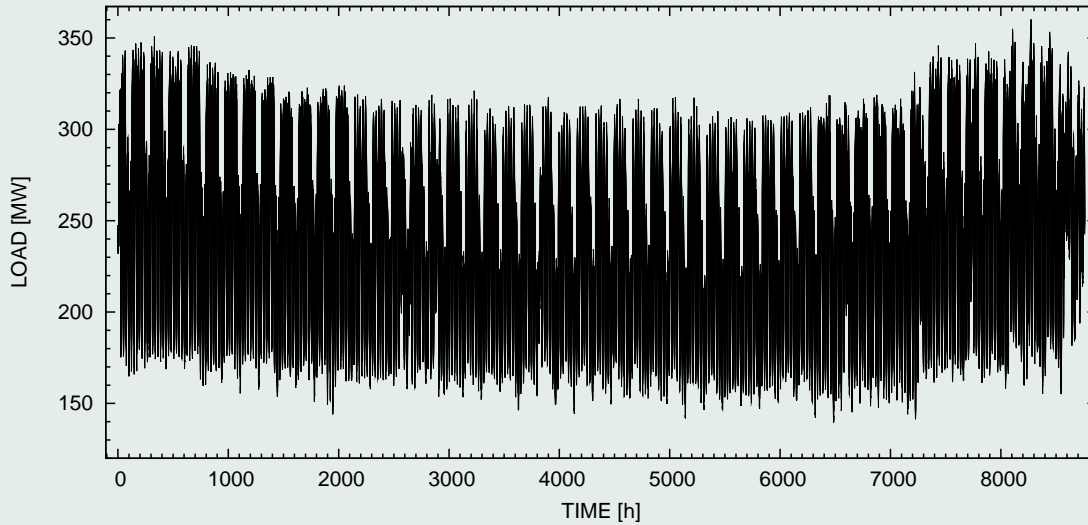
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Time plot of yearly load profile

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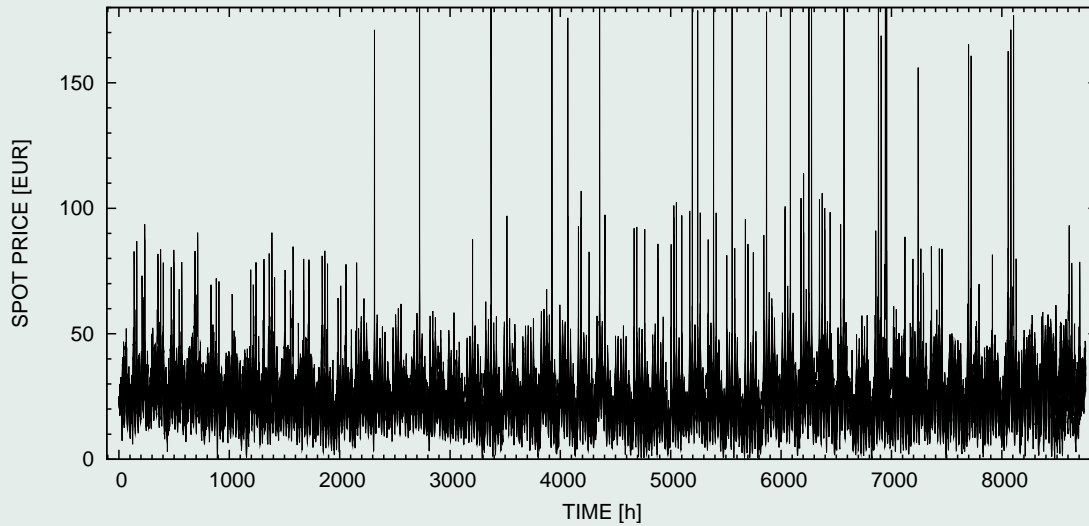
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Time plot of yearly spot price profile

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Statistical models and scenario trees

For the **stochastic input data** of the optimization model (here **yearly electricity and heat demand, and electricity spot prices**), a statistical model is employed.

- **cluster classification** for the intra-day (demand and price) profiles,
- **Three-dimensional time series model** for the daily average values (deterministic trend functions, a trivariate ARMA model for the (stationary) residual time series),
- **Generation of scenarios** by computing **Monte Carlo samples from the multivariate normal distribution** that corresponds to the ARMA process, and adding on trend functions as well as matched intra-day profiles from the clusters afterwards,

Intended modification: **QMC samples** instead of MC.

- **generation of scenario trees** (Heitsch-Römisch 09).

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Numerical results

Test runs were performed on real-life data of a German municipal power company leading to a linear program containing $T = 365 \cdot 24 = 8760$ time steps, a scenario tree with 40 demand-price scenarios (see below) with about 150.000 nodes. The objective function is of the form

$$\text{Minimize } \gamma \rho(z) - (1 - \gamma) \mathbb{E}(z_T)$$

with a (multiperiod) risk measures ρ with risk aversion parameter $\gamma \in [0, 1]$ ($\gamma = 0$ corresponds to the risk-neutral case).

Two risk measures:

(1) $\rho(z) = \text{AV@R}_{0.05}(z_T)$ (Average or Conditional Value-at-risk)

(2) $\rho(z) = \rho_m(z) = \text{AV@R}_{0.05}\left(\min_{j=1, \dots, J} z_{t_j}\right)$

($t_j, j = 1, \dots, J = 52$, are the risk measuring time steps; they correspond to 11 pm at the last trading day of each week).

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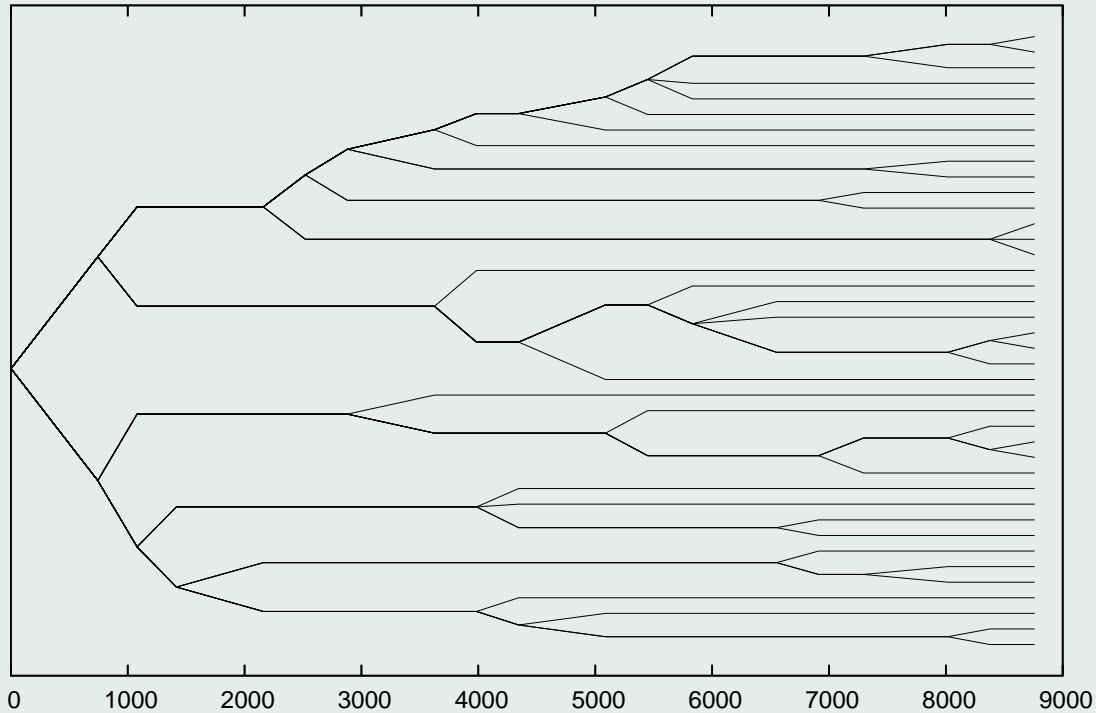
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Yearly scenario tree for the trivariate load-price process

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It turns out that the numerical results for the **expected maximal revenue and minimal risk**

$$\mathbb{E}(z_T^{*\gamma}) \quad \text{and} \quad \rho(z_{t_1}^{*\gamma}, \dots, z_{t_J}^{*\gamma})$$

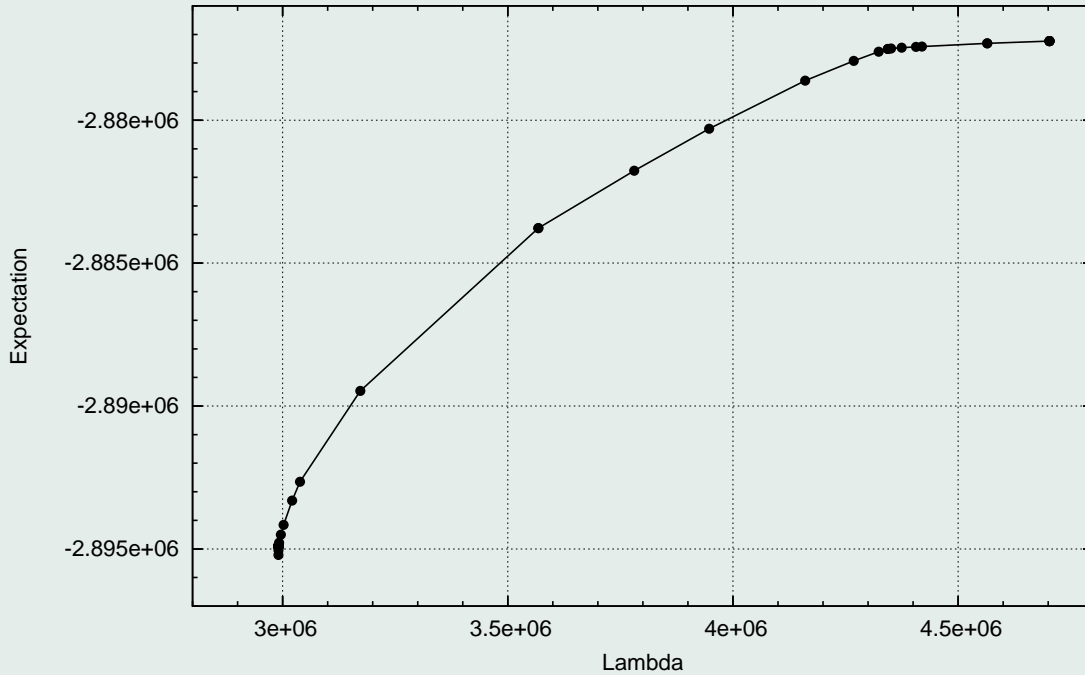
with the optimal revenue process $z^{*\gamma}$ are (almost) **identical** for $\gamma \in [0.15, 0.95]$ and the risk measures used in the test runs.

The **efficient frontier**

$$\gamma \mapsto (\rho(z_{t_1}^{*\gamma}, \dots, z_{t_J}^{*\gamma}), \mathbb{E}(z_T^{*\gamma}))$$

is **concave** for $\gamma \in [0, 1]$.

Risk aversion costs less than 1% of the expected overall revenue.



Efficient frontier

The LP is solved by CPLEX 9.1 in about 1 h running time on a 2 GHz Linux PC with 1 GB RAM.

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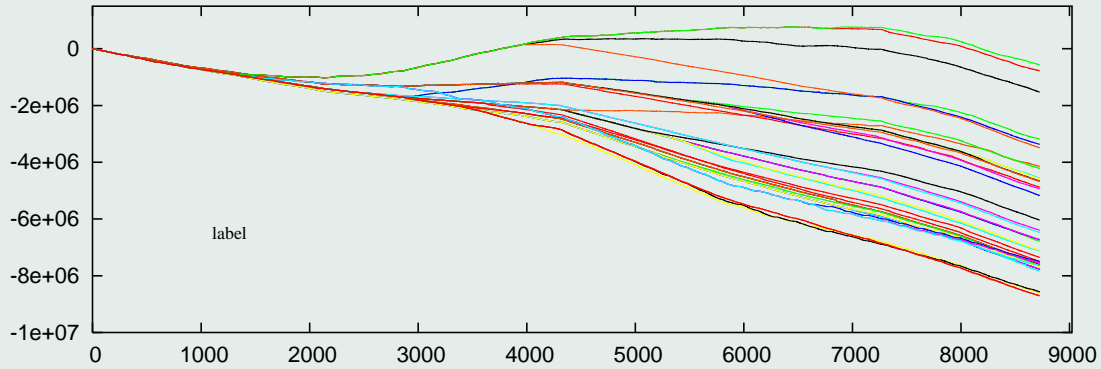
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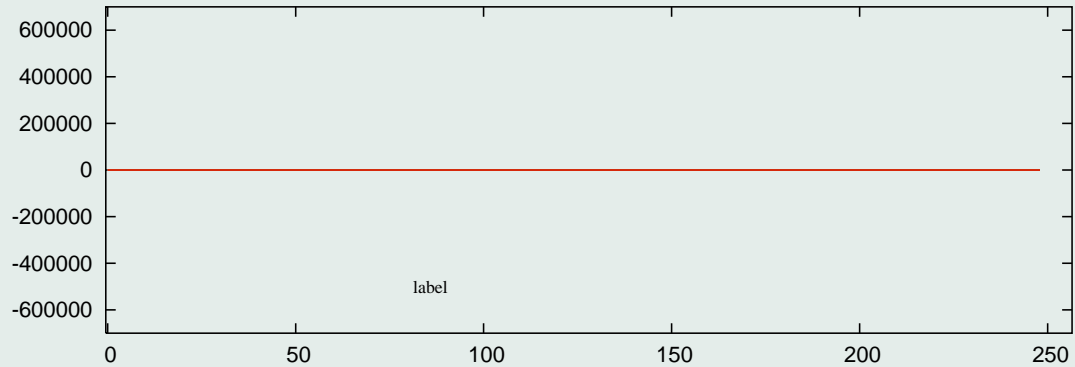
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Overall revenue scenarios for $\gamma = 0$



Future trading for $\gamma = 0$

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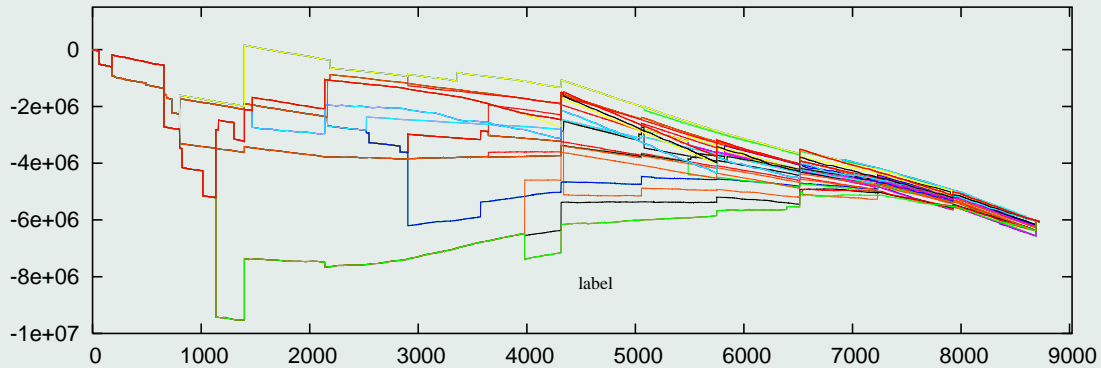
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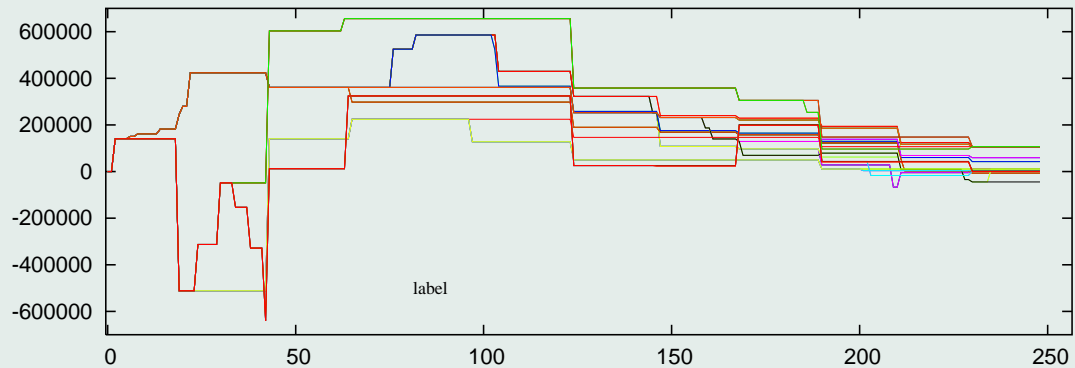
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Overall revenue scenarios with $\Delta V@R_{0.05}$ and $\gamma = 0.9$



Future trading with $\Delta V@R_{0.05}$ and $\gamma = 0.9$

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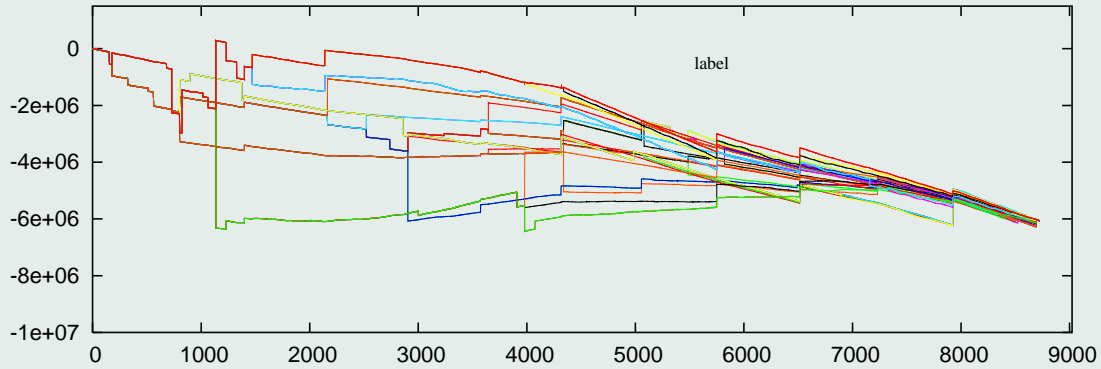
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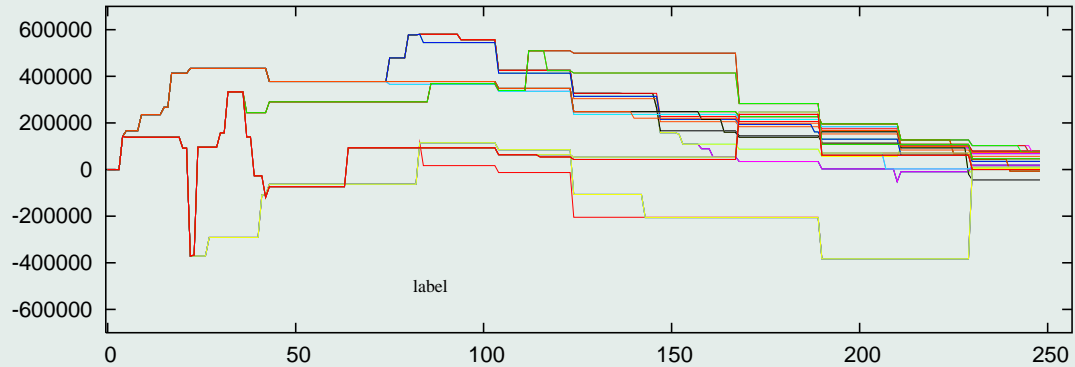
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Overall revenue scenarios with ρ_m and $\gamma = 0.9$



Future trading with ρ_m and $\gamma = 0.9$

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Conclusions

- A survey of [approaches for scenario generation in stochastic optimization](#) was presented.
- We outlined that a [theoretical basis for applying Quasi-Monte Carlo in stochastic programming](#) is still open.
- Strategies for [scenario reduction and scenario tree generation](#) were briefly discussed.
- Numerical results for a [risk-neutral and risk-averse yearly electricity portfolio management model](#) were presented.

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